# On the Moments of the Multiple Correlation Coefficient in Samples from Normal Population 

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Wishart (1930) has found the mean and the second moment coefficient of the multiple correlation coefficient in samples from a normal population when $b=1,2,3$ and then assumed that the result is probably true without any restrictions as to whether $2 b$ is even or odd. Also he has pointed out the difficulties in finding the mean and the standard deviation of the multiple correlation coefficient. Here I have found the $m$-th moment both of the square of the multiple correlation coefficient and also of the multiple correlation coefficient for unrestricted $a, b$. I have also found a method to find the mean and the second moment of the multiple correlation coefficient approximately.

Fisher's (1928) general distribution for the multiple correlation coefficient is

Hence

$$
\begin{array}{r}
d f=\frac{\mid a+b-1}{\underline{a-1} \mid \underline{\mid b-1}} \cdot\left(1-\rho^{2}\right)^{a^{4} b} F\left(a+b, a+b, a, \rho^{2} R^{2}\right) \times \\
\left(R^{2}\right)^{a-1}\left(1-R^{2}\right)^{b-1} d\left(R^{2}\right) .
\end{array}
$$

$$
\begin{aligned}
& \mu_{m}^{\prime}\left(R^{2}\right)=\frac{\mid a+b-1}{|\underline{a-1}| \underline{b-1}}\left(1-\rho^{2}\right)^{a+b} \sum_{0}^{a} \frac{(a+b)_{s}^{2}}{\mid \underline{s}(a)_{s}} \rho^{2 s} \times \\
& \int_{0}^{1} x^{a+s+m-1}(1-x)^{b-1} d x \\
& =\frac{\mid a+b-1}{\mid a-1-b-1} \cdot\left(1-\rho^{2}\right)^{a+h} \sum_{0}^{a} \frac{(a+b)^{2}}{\mid S(a)_{s}} \rho^{2 x} \times \\
& \frac{\Gamma(a+s+m) \Gamma(b)}{\Gamma(a+m+b+s)} \text { where }(a)_{s} \\
& =a(a+1) \ldots(a+s-1)
\end{aligned}
$$

and the term by term integration is permissible as the Hypergeometric function is uniformly convergent if $0 \leqslant \rho^{2}<1$.

From (11) and (12) we have

$$
\begin{aligned}
2-3 x R(x) & +\left(x^{2}-1\right) R^{2}(x)>2-3 x \cdot \frac{x^{2}+2}{x^{3}+3 x} \\
& +\left(x^{2}-1\right) \cdot \frac{x}{x^{2}+1} \cdot \frac{x^{3}+5 x}{x^{3}+6 x^{2}+3}
\end{aligned}
$$

which is easily seen to be $>0$ for sufficiently large $x>0$.
Hence

$$
[\lambda(x)-x][2 \lambda(x)-x]-1>0
$$

which shows that the third moment of $X$, about its mean which is a measure of Skewness has the same sign as $E(X)$.

## References

1. Birnbaum, Z. W. .. Annals of Math. Statistics, 1950, 21, 272-79.
2. -..... Ibid., 1942, 13, 245-46.
3. Gordan, R. D. .. lbid., 1941, 12, 364-66.
4. Kendall .. Advanced Theory of Statistics, 1, 129-30.

Hence

$$
\begin{aligned}
\mu_{m}^{\prime}\left(R^{2}\right)= & \frac{(a+m-1) \ldots a}{(a+b+m-1) \ldots(a+b)}\left(1-\rho^{2}\right)^{a+b} \\
& \quad \times{ }_{3} F_{2}\left(a+b, a+b, a+m, a, a+m+b, \rho^{2}\right)
\end{aligned}
$$

Since

$$
\begin{aligned}
& z^{\alpha_{1}+1} \cdot{ }_{3} F_{2}\left(a_{1}+1, a_{2}, a_{3}, \rho_{1}, \rho_{2}, z\right) \\
& =\frac{z^{2}}{a_{1}} \frac{d}{d_{z}} z^{\alpha_{1}}{ }_{3} F_{2}\left(a_{1}, a_{2}, a_{3}, \rho_{1}, \rho_{2}, z\right)
\end{aligned}
$$

We have

$$
\begin{gathered}
\mu_{m}^{\prime}\left(R^{2}\right)=\frac{\left(1-\dot{\rho}^{2}\right)^{a+b}}{(a+b) \ldots(a+b+m+1)} \times \\
D^{m} F\left(a+b, a+b, a+m+b, \rho^{2}\right)
\end{gathered}
$$

where

$$
D=\frac{\rho^{3}}{2} \frac{\partial}{\partial \rho}
$$

Similarly we can prove that

$$
\begin{gathered}
\mu_{m}{ }^{\prime}(R)=\frac{\mid a+b-1}{\underline{a-1}}\left(1-\rho^{2}\right)^{a+b} \frac{\Gamma\left(a+\frac{m}{2}\right)}{\Gamma\left(a+b+\frac{m}{2}\right)} \times \\
=\frac{\left(1-\rho^{2}\right)^{a+l} \Gamma\left(a+\frac{m}{2}\right)}{\Gamma\left(a+b+\frac{m}{2}\right)} D^{m} F\left(a+b+\frac{m}{2}, a+b, \frac{1}{2}, \rho^{2}\right)
\end{gathered}
$$

when $b$ is integer.
Putting

$$
\begin{gathered}
m=1 \\
\mu_{1}^{\prime}(R)=\bar{R}=\frac{\left\lvert\, a+b-1 \Gamma\left(a+\frac{1}{2}\right)\right.}{\overline{a-1} \Gamma a+b+\frac{1}{2}}\left(1-\rho^{2}\right)^{a+b} \times \\
{ }_{3} F_{2}\left(a+b, a+b, a+\frac{1}{2}, \overline{a,} a+b+\frac{1}{2}, \rho^{2}\right)
\end{gathered}
$$

when

$$
\rho=0, \bar{R}=\frac{a+b-1 \cdot \Gamma\left(a+\frac{1}{2}\right)}{a-1 \Gamma\left(a+b+\frac{1}{2}\right)}
$$

which agrees with Wishart's result.

Example.-Let us consider the example considered by Pearson (1930). Let $\rho^{2}=\cdot 5, a=3, b=47$.

$$
\begin{aligned}
& \qquad \begin{aligned}
\bar{R} & =\frac{49 \Gamma 3 \cdot 5}{2 \Gamma 50 \cdot 5}(\cdot 5)^{50}{ }_{3} F_{2}(50,50,3 \cdot 5,3,50 \cdot 5, \cdot 5) \\
& =\frac{49 \Gamma 3 \cdot 5}{2 \Gamma 50 \cdot 5}(\cdot 5)^{50}{ }_{3} F_{2}(50,50,3 \cdot 5,3 \cdot 5,51, \cdot 6) \\
= & \frac{149 \Gamma 3 \cdot 5}{2 \Gamma 50 \cdot 5}(\cdot 5)^{50} \mathrm{~F}(50,50,51, \cdot 6)=\cdot 4 \text { approximately. } \\
\quad & \text { References }
\end{aligned} \\
& \begin{array}{l}
\text { Wishart, J. } \\
\text { Fisher, R. A. } \\
\text { Pearson, K. }
\end{array} \quad \begin{array}{l}
\text {. }
\end{array} \quad \text { Biometrica, 1930, 22, } 224 .
\end{aligned}
$$

